Mortality and Deprivation

Torsten Kleinow joint work with Jie Wen and Andrew J.G. Cairns

Heriot-Watt University, Edinburgh

Actuarial Research Centre, IFoA

ARC webinar - Edinburgh - 2 October 2018









Research Project: Modelling, Measurement and Management of Longevity and Morbidity Risk

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Research Project: Modelling, Measurement and Management of Longevity and Morbidity Risk

The Research Team:

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Plus: 2 postdoctoral researchers; 3 PhD students

Plus: Aarhus, Durham, U. California.



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Index of Multiple Deprivation

The IMD is a weighted combination of seven indices of deprivation:

- Income (22.5%)
- Employment (22.5%)
- Education (13.5%)
- Health (13.5%)
- Crime (9.3%)
- Barriers to Housing and Services (9.3%)
- Living environment (9.3%)

source: GOV.UK



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- just over 30,000 LSOAs (Lower Layer Super Output Area) in England
- ordered and split into ten deciles:

10% most deprived, ..., 10% least deprived



Index of Multiple Deprivation (IMD) areas





- We consider mortality data for males in England for the ten IMD deciles (2015).
- ages: 40-89, years: 2001-2015
- source: Office for National Statistics





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- We observe different death rates at different ages, in different years and in different IMD deciles
- Death rates for most deprived are higher than death rates for least deprived.
- So, the ratio of deaths rates in most deprived areas compared to least deprived areas is greater than one,

 $\frac{\text{deaths per 1,000 lives in most deprived areas}}{\text{deaths per 1,000 lives in least deprived areas}} > 1$



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Deaths per 1,000 lives most deprived least deprived ratio 2001 25.3 11.4 2.219



male mortality in year 2001

- roughly linear in age (Gompertz line)
- mortality differentials are decreasing with age



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- similar shape as male log mortality, but lower level, slightly smaller differences
- again, mortality differentials are decreasing with age



Poll 2: Death rates by IMD decile fourteen years later

What is the ratio of death rates in the most deprived areas compared to the least deprived in England in 2015 (males, age 65)?



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	Deaths per 1,000 lives			
	most deprived	least deprived	ratio	
2001	25.3	11.4	2.219	
2015	22.3			



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2015	22.3	6.8	3.279	





male mortality in year 2015

- similar shape as in 2001
- differences at high ages are larger



- downward shift from 2001 to 2015
- differences between most deprived and least deprived have increased since 2001
- higher differences at high ages



	Deaths per 1,000 lives			
	most deprived	least deprived	ratio	
2001	25.3	11.4	2.219	
2005	27.0	9.7	2.784	
2010	23.0	8.2	2.805	
2015	22.3	6.8	3.279	





downward trend strongest for least deprived

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Period effect in Lee-Carter model by IMD decile



- downward trend strongest for least deprived
- no improvements for most deprived since 2011
- slowdown of improvements for least deprived since 2011



Focus of our research:

- Stochastic models that describe mortality experiences in all socio-economic groups simultaneously.
- Model uncertainty addressed by comparing a wide variety of models (Goodness of fit, robustness, ...)
- Leading to projections, and more importantly, **mortality scenario** generation allowing us
 - to put probabilities on certain scenarios and ...
 - then use those for Value at Risk calculations, annuity pricing, etc.



$D_{xti} \sim \text{Poisson}(m_{xti}E_{xti})$

For each period (calendar year) t, age x and IMD decile i we have

- D_{xti}: Number of deaths,
- E_{xti} : Central exposure-to-risk
- m_{xti} : force of mortality



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D_{xti}: Number of deaths,

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 m_{xti} : force of mortality

So, expected number of deaths, $E[D_{xti}] = m_{xti}E_{xti}$ Aim or our research: compare different models for the force of mortality m_{xti} .



Models

All considered models are variants of group specific Lee-Carter type models with the extension to a second age-period effect by Renshaw & Haberman (2003):

$$\log m_{xti} = \alpha_{xi} + \beta_{xi}^1 \kappa_{ti}^1 + \beta_{xi}^2 \kappa_{ti}^2 + \gamma_{ci}$$

where c = t - x is the cohort (year of birth).



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where c = t - x is the cohort (year of birth). Specific versions include models with: common age effect : $\alpha_{xi} = \alpha_x$ non-parametric common age effects : $\beta_{xi}^k = \beta_x^k$ (Kleinow, 2015) fixed age effects : constant $\beta_{xi}^1 = 1$ and linear $\beta_{xi}^2 = x - \bar{x}$, where \bar{x} is the mean age in the data set. (Plat, 2009) common period effects : $\kappa_{ti}^k = \kappa_t^k$ (Li and Lee, 2005) group specific trends in common period effects : $\kappa_{ti}^k = \kappa_t^k + \eta_i(t - \bar{t})$ and variations with and without cohort effects.



$$\log m_{xti} = \alpha_{xi} + \beta_{xi}^1 \kappa_{ti}^1 + \beta_{xi}^2 \kappa_{ti}^2 + \gamma_{ci}$$

- What parameters should be chosen to be group specific and which parameters are common?
- Should age-effects be estimated?
- Should we include cohort effects (common or group specific)?
- What parameters show the greatest differences between IMD groups?
- Are the groups clustered?



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- The improvement rates (from 2001 2015) are also different.



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- For a wider age range, models with common non-parametric age effects (Kleinow (2015) + common α) produce a good fit in terms of BIC, heatmaps ...
- However, for a narrower age range (65-89), models with constant/linear β 's, (Plat (2009) + common α) are better.



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- For a wider age range, models with common non-parametric age effects (Kleinow (2015) + common α) produce a good fit in terms of BIC, heatmaps ...
- However, for a narrower age range (65-89), models with constant/linear β 's, (Plat (2009) + common α) are better.
- Cohort effects do not improve the fit for those models
- If a cohort effect is included it should be a common cohort effect





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Critical illness insurance rates and related morbidity trends

Dr George Streftaris

Joint work with Chunxiao Xie (PhD, HWU) Dr Erengul Dodd (Southampton U) Dr Ayse Arik (HWU)

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Critical illness insurance: Policy description

- Fixed term policy, usually ceasing at age 65
- A fixed sum insured payable on the diagnosis of one of a specified list of critical illnesses
- Covers: Cancer; *Death*; Heart attack; Stroke; Multiple Sclerosis; Total & permanent disability; Coronary artery bypass graft; Kidney failure; Major organ transplant etc.
- Policies are often sold together with term or endowment insurance
- Benefit type: Full Accelerated (FA) or Stand Alone (SA)





Data Provided by the CMI Assurances Committee (UK)

- 1999-2005
 - Details of policies inforce at the start and end of each year
 - 19,127 claims settled
- 2007-2010
 - Grouped by various risk factors
 - 20,487 claims settled





Data:

- Claims
- Exposures
- Risk factors:



Risk factor (covariate)	1999 – 2005	2007 – 2010
Age (last birthday)	\checkmark	\checkmark
Gender	\checkmark	\checkmark
Smoker		\checkmark
Policy duration	\checkmark	\checkmark
Office	\checkmark	
Distribution channel	\checkmark	\checkmark
Benefit type (accelerated, standalone)	\checkmark	
Benefit amount	\checkmark	\checkmark
Policy type (single, joint)		
Settlement year	\checkmark	\checkmark
Cause		
Product category	\checkmark	\checkmark
Date of diagnosis	\checkmark	

Data: 2007 - 2010







31.4%

18.3%

0.1%



Single-tied

Unknown







Claims edata: 2007 - 2010 v 1999 - 2005



- Distributions very similar between 2007 2010 & 1999 2005
- Slightly higher proportion of **F** and **NS** in 2007 2010
- Lower proportion of age 16-30 in 2007 2010







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UK CII claim rates in 2007-2010 (as compared to 1999 – 2005):

(a) have gone considerably up;
(b) have gone considerably down;
(c) have stayed roughly unchanged;
(d) I don't know.

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Modelling: mostly Bayesian stochastic

- Estimation & smoothing of CI diagnosis rates
 - how do these depend on risk factors?
- Diagnosis is the insured event and there is a delay between diagnosis and settlement



- For 1999- 2005 data:
 - exposure corresponds to claims settled, not to claims diagnosed
 - we have made adjustments by fitting a delay distribution (Bayesian Generalised Beta 2 model)





Stochastic modelling: Claim rates



λ^(j)_{x,θ} : diagnosis (claim) rate for cause *j* at age *x* with risk factors θ

 $\lambda_{x,\theta} \sim LN(\delta x + \beta \theta, \sigma^2)$

• normal priors for coefficient vectors δ and β .





Stochastic modelling: Risk factor estimates for claim rates (2007 – 2010)



Risk factors: Bayesian estimates

HERIOT WATT UNIVERSITY Perform variable (factor) selection

Selected model includes:

- ✓ age (older ↑)
- ✓ smoker status (S ↑)
- ✓ distribution channel
- ✓ benefit type (stand-alone ↓)
- ✓ age x smoker

Also (not shown here):

✓ policy duration (longer ↑)
✓ benefit amount (mid ↑)



Fitted claim rates (and intervals) 2007-2010 v 1999 – 2005, Accelerated, Smoker, Pol Duration 1



- Model fits crude rates (2007 2010) well
- 2007 2010 rates significantly higher
- Gap wider at younger ages
- Similar trends for other profiles



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Fitted claim rates

Smokers & non-smokers (Accelerated, Pol Duration 1)



Fitted claim rates Accelerated v Stand alone (2007 – 2010) & 1999 – 2005

Inception Rate for Different Benefit type Smoker with PolDur 4



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- ✤ Accelerated 2007 -2010 (black) higher than stand-alone (green)
- Both significantly higher than 1999 - 2005



UK population cancer rates (ONS data) All cancers

52

57 62



Fitted: ——; Observed: •••

- Bayesian GLM with: age, year, gender
- Incidence rates increasing with time
- Higher rates for older ages





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Insured population cancer rates in 1999-2005 (as compared to general population cancer rates):

(a) are at the same level;
(b) are considerably higher;
(c) are considerably lower;
(d) I don't know.

Population cancer rates v insurance rates Males - All cancers





- Experience for the insured population is different
- CII rates significantly lower than population rates

✤ Why?

- -- Differences between those who can/cannot afford CII?
- -- Rates lower in most affluent groups? (but not for all types of cancer)
- -- Underwriting effect?





Population cancer rates v insurance rates Females - All cancers

52

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- Gap smaller than for males (for older ages)
- Effect of breast cancer? (same for all socio-economic groups)



Population cancer rates v insurance rates Females – Excluding melanoma skin cancer

62



- Some cancers not covered by CII
- Exclude skin cancer from population rates:
 - gap now smaller
 - CII rates increasing faster that population rates?





Summary

- CII claimants distribution similar between 1999-2005 & 2007-2010
- Claim rates (2007-2010) depend on a number of risk factors:
 - age, smoker status, distribution channel, policy duration, benefit amount and benefit type, etc.
- Analysis suggests increase of CII claim rates over time
 - especially at younger ages
- Cancer: insurance rates much lower than population rates
- But trends could be different (worse for CII)?







The views expressed in this presentation are those of the presenter.

