# MULTI-POPULATION MORTALITY MODELLING Danish data and Basis Risk

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#### Plan

- Motivation and challenges
- Danish males data
  - 10 sub-populations grouped by affluence
- An extended CBD multi-population model
- Bayesian implementation and results

### Motivation for multi-population modelling

#### A: Risk assessment

- Multi-country (e.g. consistent demographic projections)
- Males/Females (e.g. consistent demographic projections)
- Socio-economic subgroups (e.g. blue or white collar)
- Smokers/Non-smokers
- Annuities/Life insurance
- Limited data ⇒ learn from other populations

#### Motivation for multi-population modelling

## B: Risk management for pension plans and insurers

- Retain systematic mortality risk; versus:
- 'Over-the-counter' deals (e.g. longevity swap)
- Standardised mortality-linked securities
  - linked to national mortality index
  - < 100% risk reduction: basis risk

## Challenges

- Data availability
- Data quality and depth
- Model complexity
  - single population models can be complex
  - 2-population versions are more complex
  - multi-pop .....
- Multi-population modelling requires
  - (fairly) simple single-population models
  - simple dependencies between populations

## A New Case Study and a New Model

- Sub-populations differ from national population
  - socio-economic factors
  - geographical variation
  - other factors
- Denmark
  - High quality data on ALL residents
  - 1981-2005 available
  - Can subdivide population using covariates on the database

#### **Danish Data**

 What can we learn from Danish data that will inform us about other populations?

- Key covariates
  - Wealth
  - Income
- Affluence = Wealth $+15 \times$  Income

#### **Problem**

- ◆ High income ⇒ "affluent" and healthy BUT
- Low income 

  → not affluent, poor health
- ◆ High wealth ⇒ "affluent" and healthy BUT

#### Solution: use a combination

- ullet Affluence, A= wealth  $+K\times$  income
- $\bullet$  K=15 seems to work well *statistically* as a predictor
- $\bullet$  Low affluence, A, predicts poor mortality

## Subdividing Data

- Males resident in Denmark for the previous 12 months
- Divide population in year t
  - into 10 equal sized Groups (approx)
  - using *affluence*, A
- Individuals can change groups up to age 67
- Group allocations are locked down at age 67

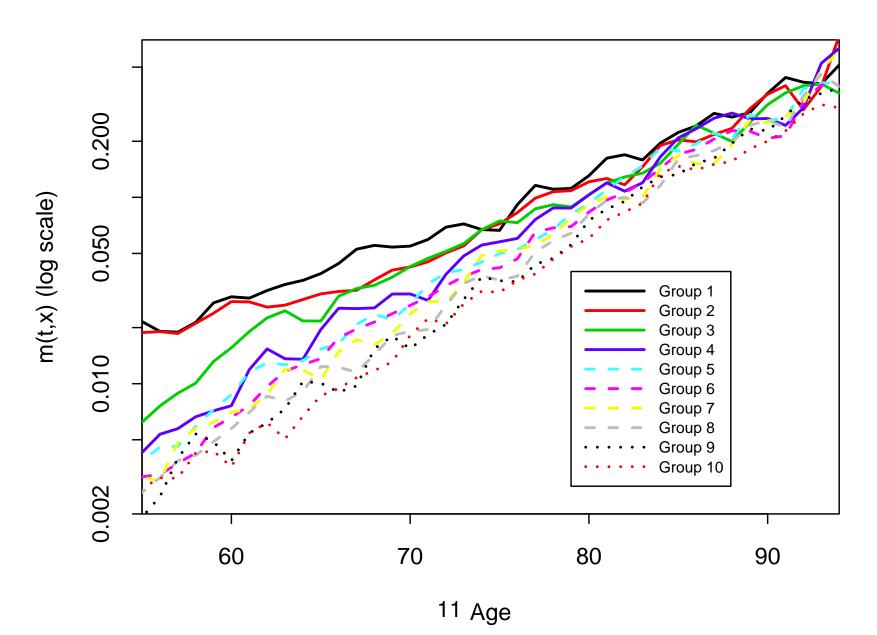
(better than not locking down at age 67)

#### Subdivided Data

- ullet Exposures  $E^{(i)}(t,x)$  for groups  $i=1,\ldots,10$  range from over 4000 down to 20
- $\bullet$  Deaths  $D^{(i)}(t,x)$  range from 150 down to 6
- $\bullet$  Crude death rates  $\hat{m}^{(i)}(t,x) = D^{(i)}(t,x)/E^{(i)}(t,x)$
- Small groups ⇒ Poisson risk is important

#### Crude death rates 2005

#### Males Crude m(t,x); 2005



## Modelling the death rates, $m_k(t,x)$

 $m^{(k)}(t,x) = \text{pop. } k \text{ death rate in year } t \text{ at age } x$ 

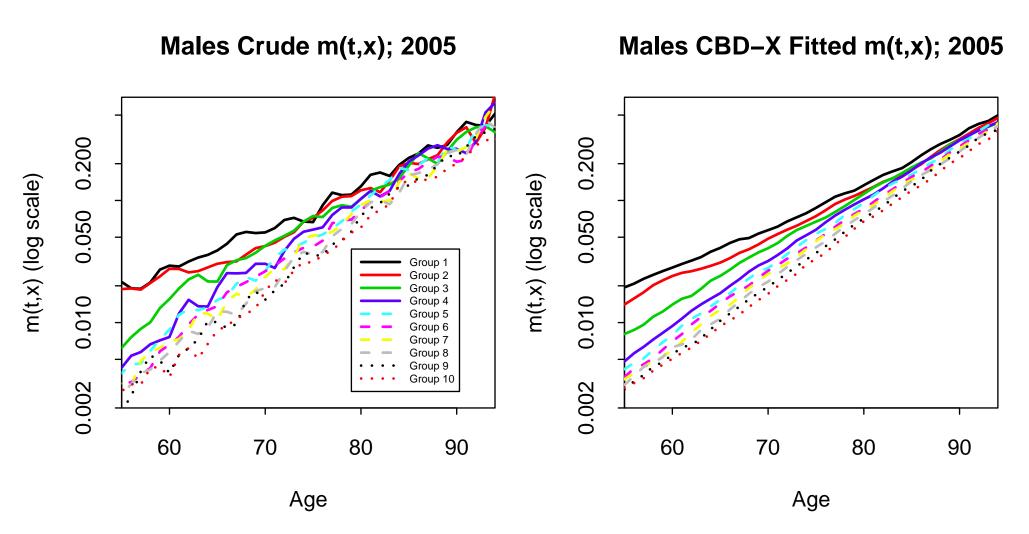
Population k, year t, age x

$$\log m^{(k)}(t,x) = \beta^{(k)}(x) + \kappa_1^{(k)}(t) + \kappa_2^{(k)}(t)(x - \bar{x})$$

(Extended CBD with a non-parametric base table,  $\beta^{(k)}(x)$ )

- 10 groups,  $k=1,\ldots,10$  (low to high affluence)
- 21 years,  $t = 1985, \dots, 2005$
- 40 ages,  $x = 55, \dots, 94$

## Model-Inferred Underlying Death Rates 2005



## Modelling the death rates, $m_k(t,x)$

$$\log m^{(k)}(t,x) = \beta^{(k)}(x) + \kappa_1^{(k)}(t) + \kappa_2^{(k)}(t)(x - \bar{x})$$

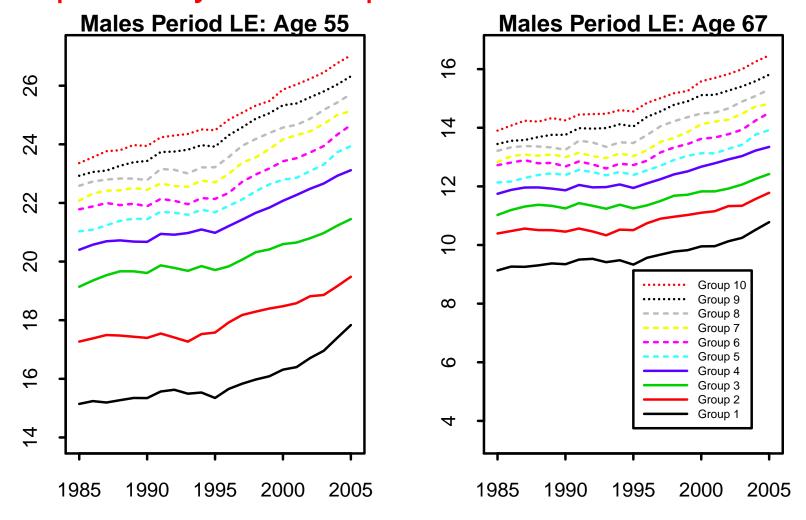
- Model fits the 10 groups well without a cohort effect
- $\bullet$  Non-parametric  $\beta^{(k)}(x)$  is essential to preserve group rankings
  - Rankings are evident in crude data
  - "Bio-demographical reasonableness":

more affluent  $\Rightarrow$  healthier

## Time series modelling

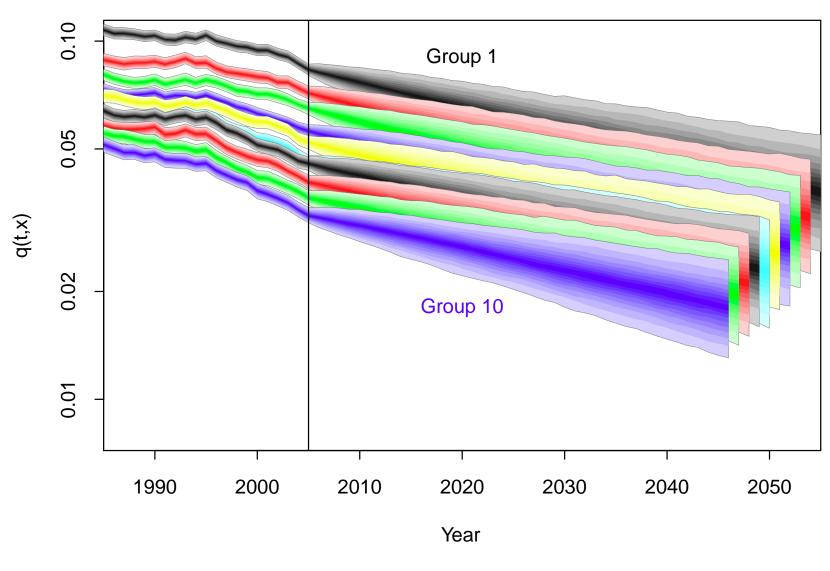
- $t \rightarrow t + 1$ : Allow for correlation
  - between  $\kappa_1^{(k)}(t+1)$  and  $\kappa_2^{(k)}(t+1)$
  - between groups  $k=1,\ldots,10$
- Bio-demographical reasonableness
  - ⇒ key hypothesis: groups should not diverge
  - ⇒ group specific period effects gravitate towards the national trend

## Life Expectancy for Groups 1 to 10

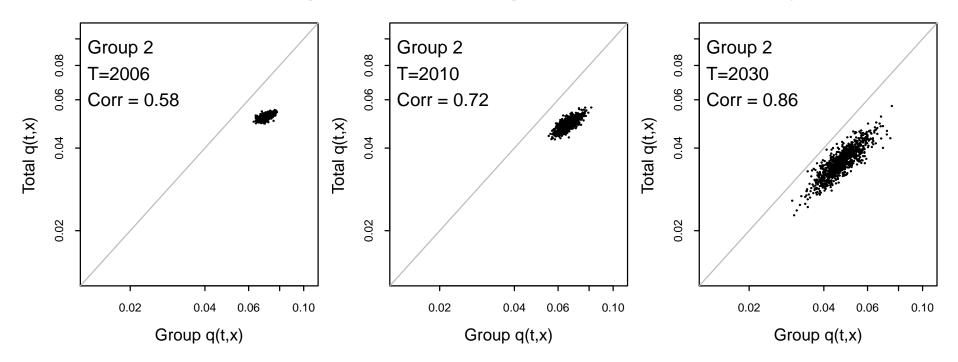


## Mortality Fan Charts Including Parameter Uncertainty

**Mortality Rates: Age 75** 



## Simulated Group versus Population Mortality



#### As T increases

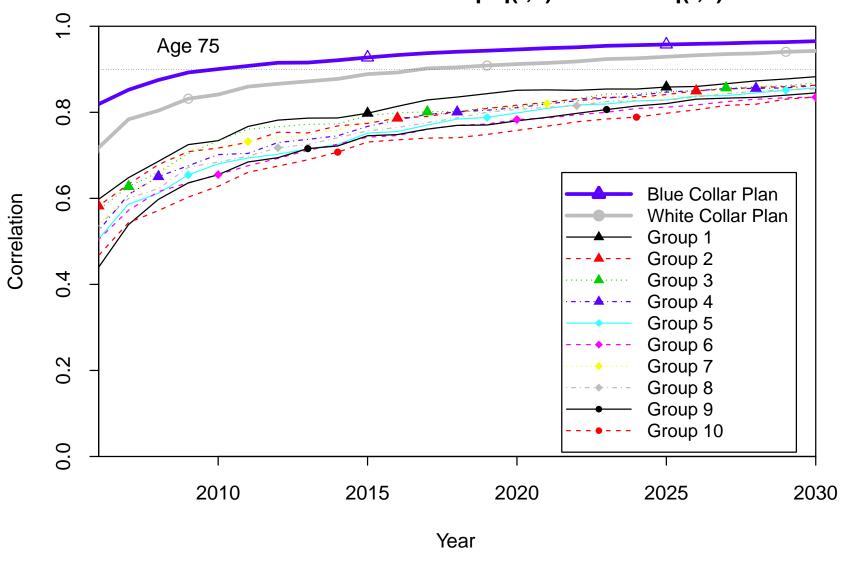
- Scatterplots become more dispersed
- Shift down and to the left
- Correlation increasess

#### **Forecast Correlations**

- Deciles are quite narrow subgroups
- Blue collar pension plan
  - $\Rightarrow$  equal proportions of groups 2, 3, 4
- White collar pension plan
  - $\Rightarrow$  equal proportions of groups 8, 9, 10

#### **Forecast Correlations**

#### Correlation Between Group q(t,x) and Total q(t,x)



#### Conclusions

- Development of a new multi-population dataset for Denmark strong bio-demographically reasonable group rankings based on a new measure of affluence
- Unlike multi-country data
   a priori ranking of affluence-related groups
- Proposal for a simple new multi-population model
- Strong correlations over medium to long term
- Correlations depend strongly on diversity of sub-population

## **Bonus Slides**

#### A specific model

$$\begin{array}{lll} \kappa_1^{(i)}(t) &=& \kappa_1^{(i)}(t-1) + \mu_1 + Z_{1i}(t) & \text{(random walk)} \\ && -\psi\left(\kappa_1^{(i)}(t-1) - \bar{\kappa}_1(t-1)\right) & \text{(gravity between groups)} \\ \kappa_2^{(i)}(t) &=& \kappa_2^{(i)}(t-1) + \mu_2 + Z_{2i}(t) \\ && -\psi\left(\kappa_2^{(i)}(t-1) - \bar{\kappa}_2(t-1)\right) \end{array}$$

where

$$ar{\kappa}_1(t) = rac{1}{n} \sum_{i=1}^n \kappa_1^{(i)}(t)$$
 and  $ar{\kappa}_2(t) = rac{1}{n} \sum_{i=1}^n \kappa_2^{(i)}(t)$ 

#### A specific model

$$\kappa_1^{(i)}(t) = \kappa_1^{(i)}(t-1) + \mu_1 + Z_{1i}(t) - \psi \left( \kappa_1^{(i)}(t-1) - \bar{\kappa}_1(t-1) \right) 
\kappa_2^{(i)}(t) = \kappa_2^{(i)}(t-1) + \mu_2 + Z_{2i}(t) - \psi \left( \kappa_2^{(i)}(t-1) - \bar{\kappa}_2(t-1) \right)$$

#### Model structure $\Rightarrow$

- $(\bar{\kappa}_1(t), \bar{\kappa}_2(t)) \sim$  bivariate random walk
- $\bullet$  Each  $\kappa_1^{(i)}(t) \bar{\kappa}_1(t) \sim AR(1)$  reverting to 0
- $\bullet$  Each  $\kappa_2^{(i)}(t) \bar{\kappa}_2(t) \sim AR(1)$  reverting to 0
- $\beta^{(i)}(x)$  vs  $\beta^{(j)}(x)$   $\Rightarrow$  intrinsic group differences

## Non-trivial correlation structure: between different ages and groups

$$\kappa_1^{(i)}(t) = \kappa_1^{(i)}(t-1) + \mu_1 + Z_{1i}(t) - \psi \left( \kappa_1^{(i)}(t-1) - \bar{\kappa}_1(t-1) \right) 
\kappa_2^{(i)}(t) = \kappa_2^{(i)}(t-1) + \mu_2 + Z_{2i}(t) - \psi \left( \kappa_2^{(i)}(t-1) - \bar{\kappa}_2(t-1) \right)$$

The  $Z_{ki}$  are multivariate normal, mean 0 and

$$Cov(Z_{ki}, Z_{lj}) = \begin{cases} v_{kl} & \text{for } i = j \\ \rho v_{kl} & \text{for } i \neq j \end{cases}$$

 ${\color{blue} \rho}=$  cond. correlation between  $\kappa_1^{(i)}(t)$  and  $\kappa_1^{(j)}(t)$  etc.

#### Comments

- Model is very simple
  - One gravity parameter,  $0<\psi<1$
  - One between-group correlation parameter,

$$0 < \rho < 1$$

- Many generalisations are possible
- But more parameters + more complex computing
- This simple model seems to fit quite well.
- Nevertheless ⇒ work in progress

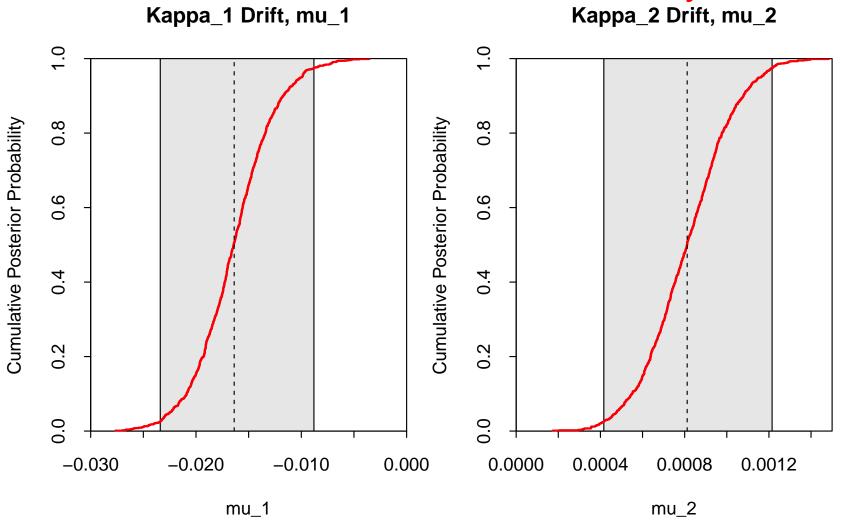
#### Prior distributions

- As uninformative as possible
- $\mu_1, \; \mu_2 \sim$  improper uniform prior
- $\{v_{ij}\}$  ~ Inverse Wishart
- $\bullet \ \rho \sim \text{Beta}(2,2)$
- $\bullet \ \psi \sim \mathrm{Beta}(2,2)$

State variables and process parameters estimated using

MCMC (Gibbs + Metropolis-Hastings)

## Posterior Distributions and 95% Credibility Intervals



Note:  $-\mu_1$  = global improvement rate

## Posterior Distributions and 95% Credibility Intervals

