# THE USE OF AUXILIARY INFORMATION TO DEAL WITH INFORMATIVELY MISSING OR OBSERVED DATA 

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## 'MISSING' OBSERVATIONS

- Response variable of interest $Y$ :
- Observed: $Y^{o}$
- Missing: $Y^{m}$
- Explanatory variable(s) or covariate(s): $X$
- Response or observation indicator: $R$


## 'MISSING' OBSERVATIONS

- Missing Completely at Random (MCAR)
$-f\left(R \mid Y^{o}, Y^{m}, X\right)=f(R)$
- Covariate Dependent-MCAR (CD-MCAR)
$-f\left(R \mid Y^{o}, Y^{m}, X\right)=f(R \mid X)$
- Covariate Dependent Missing at Random (CD-MAR)
$-f\left(R \mid Y^{o}, Y^{m}, X\right)=f\left(R \mid X, Y^{o}\right)$
- Missing Not at Random (MNAR)

$$
-f\left(R \mid Y^{o}, Y^{m}, X\right) \neq f\left(R \mid X, Y^{o}\right)
$$

## Two Modelling Approaches for MNAR Data

- Selection Model:
$f(Y, R)=f(Y) f(R \mid Y)$
or
$f(Y, R \mid X)=f(Y \mid X) f(R \mid Y, X)$ conditioning on X
- Pattern Mixture Model:

$$
f(Y, R)=f(Y \mid R) f(R)
$$

or
$f(Y, R \mid X)=f(Y \mid R, X) f(R \mid X)$ conditioning on X

## National Survey of Sexual Attitudes and Lifestyles (NATSAL)

Involved face-to-face questioning and a self-completion booklet with more sensitive questions.

- Responders: provided answers to all questions
- Item non-responders: refused to answer some questions
- Unit non-responders: refused to answer any questions

Mock Example: Level of Virginity

$\left.$|  | Responders | Item | Non-responders |
| :--- | :---: | :---: | :---: | | Unit |
| :---: |
| Non-responders | \right\rvert\,

Estimate of level of virginity
Responders only: 12.5\%

Mock Example: Level of Virginity

|  |  | Item | Unit |
| :---: | :---: | :---: | :---: |
|  | Responders | Non-responders | Non-responders |
| Embarrassed | 150 | 75 |  |
|  | $20 \%$ |  |  |
| Not | 450 | 25 |  |
| Embarrassed | $10 \%$ |  | 300 |
| Total | 600 | 100 |  |

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Estimate of level of virginity
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$12.5 \%$ Responders + Item-nonresponders: 13.2\%

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|  |  | Item | Unit |
| :---: | :---: | :---: | :---: |
|  | Responders | Non-responders | Non-responders |
| Embarrassed | 150 | 75 | 225 |
|  | $20 \%$ | $20 \%$ |  |
| Not | 450 | 25 | 75 |
| Embarrassed | $10 \%$ | $10 \%$ |  |
| Total | $\mathbf{6 0 0}$ | 100 | $\mathbf{3 0 0}$ |

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Mock Example: Level of Virginity

|  |  | Item | Unit |
| :---: | :---: | :---: | :---: |
|  | Responders | Non-responders | Non-responders |$|$| Embarrassed | $\mathbf{1 5 0}$ | $\mathbf{7 5}$ | $20 \%$ |
| :---: | :---: | :---: | :---: |
|  | $20 \%$ | $20 \%$ | 75 |
| Not | 450 | 25 | $10 \%$ |
| Embarrassed | $10 \%$ | $10 \%$ | $\mathbf{3 0 0}$ |
| Total | $\mathbf{6 0 0}$ | $\mathbf{1 0 0}$ |  |

Estimate of level of virginity

Responders only:
$12.5 \%$
Responders + Item-nonresponders: 13.2\%
Responders + Item-nonresponders + 14.5\% Unit-nonresponders

## Hepatitis C disease progression

- The Trent hepatitis C cohort follows patients sporadically through visits to hospital clinic
- For patients who attend clinic and are not lost to follow-up
- Liver function tests (LFTs) are blood tests collected regularly.
- Liver biopsies (invasive procedure) are infrequent and irregular. Each biopsy scored for stage of disease, e.g. $1=$ Mild, $2=$ Moderate, $3=$ Cirrhosis
- Other data collected at clinic visits (alcohol use, treatment regimes, BMI, end-stage liver diseases)


## Process of Interest <br> Progress through biopsy states



Figure 1: Fibrosis Model

## The problem

- Liver biopsies are the gold standard in assessing disease stage
- The occurrence of liver biopsies may be informative
- We need to jointly model the examination (liver biopsy) process and outcome process to obtain correct inferences.
- Can the much more frequently recorded LFTs help?

Informative examination scheme as a missing data problem

- Consider whether a biopsy has occurred in each six-month period.
- Associate a single LFT value with each six-month period.



## Notation

Observations denoted by $i=1, \ldots, n$
$Y_{i}$ - categorical outcome at time $t_{j}$ (e.g. Stage of HCV disease)

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$X_{i}$ - surrogate variable for outcome $Y_{i}$ at time $t_{i}$
$Y^{o}$ - vector of observed outcomes
$\boldsymbol{Y}^{m}$ - vector of missing outcomes

## Approaches to joint modelling of $Y$ and $R$

- There are identifiability problems in estimating relationship between $Y$ and $R$, since $Y$ is unobserved when $R=0$
- Assumptions MUST be made before carrying out any missing data analysis:

1. A covariate dependent missing at random (CD-MAR) assumption $f\left(\boldsymbol{R} \mid \boldsymbol{Y}^{o}, \boldsymbol{Y}^{m}, \boldsymbol{Z}\right)=f\left(\boldsymbol{R} \mid \boldsymbol{Y}^{o}, \boldsymbol{Z}\right)$.
If truly CD-MAR, then unbiased inferences can obtained using the observed data, and ignoring the missingness mechanism.
2. If not willing to assume CD-MAR given $Y^{0}$ and $Z$, must seek some extra information, $X$ so that $f\left(\boldsymbol{R} \mid \boldsymbol{Y}^{o}, \boldsymbol{Y}^{m}, \boldsymbol{Z}, \boldsymbol{X}\right)=$ $f\left(\boldsymbol{R} \mid \boldsymbol{Y}^{o}, \boldsymbol{Z}, \boldsymbol{X}\right)$

The Partially Hidden Markov model (PHMM)


## The Partially Hidden Markov model Likelihood

The likelihood under this model is of the following form:

$$
\propto \prod_{i=1}^{n} f\left(\boldsymbol{r}_{i} \mid \boldsymbol{x}_{i}, \boldsymbol{y}_{i}^{o}, \boldsymbol{z}_{i}, \psi\right) \sum_{\boldsymbol{y}_{i}^{m}} f\left(\left\{\boldsymbol{y}_{i}^{m}, \boldsymbol{y}_{i}^{o}\right\} \mid \boldsymbol{z}_{i}, \theta\right) f\left(\boldsymbol{x}_{i} \mid\left\{\boldsymbol{y}_{i}^{m}, \boldsymbol{y}_{i}^{o}\right\}, \boldsymbol{z}_{i}, \phi\right)
$$

where $\psi$ and $\phi$ denote the parameters defining the probability density functions $f\left(\boldsymbol{r}_{i} \mid \boldsymbol{x}_{i}, \boldsymbol{y}_{i}^{o}, \boldsymbol{z}_{i}\right)$ and $f\left(\boldsymbol{x}_{i} \mid\left\{\boldsymbol{y}_{i}^{m}, \boldsymbol{y}_{i}^{o}\right\}, \boldsymbol{z}_{i}\right)$, respectively.

- Generalization (slightly) of CD-MAR since $X$ cannot be regarded as a covariate for the $Y$ process.
- Might be termed Surrogate-Dependent MAR


## Simulation Study

- Simulation study of samples of 300 individuals observed at 5 time points.
- Exponential two-stage model, normally distributed auxiliary variable.
- Negative biases of $8 \%$ to $26 \%$ in estimation of baseline hazard if MCAR assumption is made incorrectly.
- Negative biases of $3 \%$ to $8 \%$ for a binary covariate ( $50 \%$ at each level) coefficient.
- PHMM eliminates these biases and gives appropriate coverage etc if observation depends on the auxiliary variable $X$.
- PHMM offers significant improvement even if MNAR model is correct.


## Simulation Study

- Data are generated for 300 individuals at five equally spaced examination times, $\left(t_{0}, t_{1}, t_{2}, t_{3}, t_{4}\right)=(0,2,4,6,8)$
- $Z=-1$ for $50 \%$ of individuals and 1 otherwise.
- Transition time out of state 1 is exponential
$-T \mid z \sim \operatorname{Exponential}\left(\lambda_{0} e^{\beta z}\right), \lambda_{0}=0.2, \beta=0.5$.
- The binary response $Y\left(t_{j} \mid T\right)=\mathbf{I}\left[T \leq t_{j}\right]$ indicator for transition by $t_{j}$.
- The auxiliary variables, $X$, are normally distributed
$-\left(X\left(t_{j}\right) \mid y\left(t_{j}\right)\right) \sim \operatorname{Normal}\left(\mu_{y\left(t_{j}\right)}, \sigma^{2}\right)$
$-\mu_{y\left(t_{j}\right)}=\phi_{0}+\phi_{1} y\left(t_{j}\right), \sigma=1$ (independent of $Y$ )
- Missing data process is Bernoulli

$$
-\operatorname{Pr}\left(R\left(t_{j}\right)=1 \mid y\left(t_{j}\right), x\left(t_{j}\right)\right)=\operatorname{logit}^{-1}\left\{\psi_{0}+\psi_{1} y\left(t_{j}\right)+\psi_{2} x\left(t_{j}\right)\right\} .
$$

## Simulation results for the baseline log hazard

| Scenario | Relative bias (\%) |  |  | 95\% coverage (\%) |  |  | MSE |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | IG | PHMM | MNAR | IG | PHMM | MNAR | IG | PHMM | MNAR |
| MCAR, $\psi_{1}=\psi_{2}=0$ |  |  |  |  |  |  |  |  |  |
| 1) $X$ independent of $Y, \phi_{1}=0$ | 0.1 | 0.1 | 0.1 | 95.9 | 96.0 | 94.5 | 0.009 | 0.009 | 0.013 |
| 2) $\phi_{1}=0.5$ | 0.1 | 0.1 | 0.1 | 95.9 | 95.4 | 94.5 | 0.009 | 0.008 | 0.013 |
| 3) $\phi_{1}=1$ | 0.1 | 0.1 | 0.1 | 95.9 | 94.7 | 94.5 | 0.009 | 0.008 | 0.013 |
| MAR, $\psi_{1}=0, \psi_{2}=1$ 1) $X$ independent of $Y, \phi_{1}=0$ |  |  |  |  |  |  |  |  |  |
| 1) $X$ independent of $Y, \phi_{1}=0$ | 0.2 | 0.2 | 0.2 | 96.1 | 95.7 | 95.5 | 0.008 | 0.008 | 0.011 |
| 2) $\phi_{1}=0.5$ | -8.4 | 0.2 | 0.2 | 64.2 | 96.0 | 95.2 | 0.026 | 0.008 | 0.011 |
| 3) $\phi_{1}=1$ | -16.8 | 0.2 | 0 | 12.0 | 95.0 | 95.1 | 0.081 | 0.007 | 0.011 |
| MNAR, $\psi_{1}=\psi_{2}=1$ |  |  |  |  |  |  |  |  |  |
| 1) $X$ independent of $Y, \phi_{1}=0$ | -13.5 | -14.0 | 0.0 | 21.8 | 19.1 | 96.0 | 0.053 | 0.057 | 0.008 |
| 2) $\phi_{1}=0.5$ | -19.9 | -13.1 | 0.1 | 1.1 | 25.0 | 95.4 | 0.109 | 0.051 | 0.008 |
| 3) $\phi_{1}=1$ | -26.4 | -9.9 | 0.0 | 0.0 | 48.7 | 94.1 | 0.188 | 0.032 | 0.008 |

## Simulation results for the binary covariate coefficient

| Scenario | Relative bias (\%) |  |  | 95\% coverage (\%) |  |  | MSE |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | IG | PHMM | MNAR | IG | PHMM | MNAR | IG | PHMM | MNAR |
| MCAR, $\psi_{1}=\psi_{2}=0$ |  |  |  |  |  |  |  |  |  |
| 1) $X$ independent of $Y, \phi_{1}=0$ | 0.4 | 0.3 | 0.4 | 94.8 | 94.8 | 94.9 | 0.009 | 0.009 | 0.009 |
| 2) $\phi_{1}=0.5$ | 0.4 | 0.2 | 0.4 | 94.8 | 95.4 | 94.9 | 0.009 | 0.008 | 0.009 |
| 3) $\phi_{1}=1$ | 0.4 | 0.1 | 0.4 | 94.8 | 95.0 | 94.9 | 0.009 | 0.007 | 0.009 |
| MAR, $\psi_{1}=0, \psi_{2}=1$ |  |  |  |  |  |  |  |  |  |
| 1) $X$ independent of $Y, \phi_{1}=0$ | 0.4 | 0.3 | 0.3 | 94.5 | 94.4 | 94.3 | 0.008 | 0.008 | 0.008 |
| 2) $\phi_{1}=0.5$ | -3.0 | 0.2 | 0.3 | 93.4 | 94.5 | 94.4 | 0.008 | 0.008 | 0.008 |
| 3) $\phi_{1}=1$ | -5.8 | 0.4 | 0.6 | 91.9 | 94.7 | 94.3 | 0.009 | 0.007 | 0.008 |
| MNAR, $\psi_{1}=\psi_{2}=1$ |  |  |  |  |  |  |  |  |  |
| 1) $X$ independent of $Y, \phi_{1}=0$ | -4.2 | -4.5 | 0.2 | 92.4 | 92.3 | 94.8 | 0.007 | 0.007 | 0.007 |
| 2) $\phi_{1}=0.5$ | -6.1 | -1.8 | 0.3 | 90.6 | 93.6 | 94.4 | 0.008 | 0.007 | 0.007 |
| 3) $\phi_{1}=1$ | -7.8 | 0.6 | 0.4 | 90.0 | 94.9 | 95.4 | 0.008 | 0.006 | 0.006 |

## Transitions in Trent Cohort Database

|  | To state |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| From state | 'None/Mild' 'Moderate' | 'Severe/Cirrhosis' | Unknown |  |
| 'None/Mild' | 326 | 20 | 6 | 403 |
| 'Moderate' | 0 | 8 | 6 | 109 |
| 'Severe/Cirrhosis' | 0 | 0 | 2 | 100 |

Observed disease state transitions in the Trent hepatitis C cohort.

## Basline Hazards Estimates

| Parameter | Model <br> Ignorable |  |  |
| :--- | :---: | :---: | :---: |
| CD-MAR (ALT) | MNAR |  |  |
| Baseline intensities | 0.0120 | 0.0119 | 0.0119 |
| $\lambda_{1,2}$ | $(0.0079,0.0182)$ | $(0.0078,0.0181)$ | $(0.0078,0.0181)$ |
| $\lambda_{2,3}$ | 0.0773 | 0.0769 | 0.0794 |
|  | $(0.0396,0.1509)$ | $(0.0399,0.1485)$ | $(0.0386,0.1634)$ |

## A Different Sort of Example

Return to NATSAL (Survey of Sexual Attitudes and Lifestyle)

- Two surveys in 1990 and 2000
- Interested in changes between 1990 and 2000
- As seen before, bias is expected in each survey.
- Change in bias is relevant to any examination of change in results of the surveys


## Comparison of NATSAL-1990 with NATSAL-2000

- Bias will depend on the question.
- Classify questions to be of high, medium and low sensitivity (effectively reflecting expected bias).
- Should any information be the same in the two surveys?
- Population cohort eligible for both surveys are those:
- Aged 16-34 in 1990.
- Aged 26-44 in 2000.
- Questions answered by this common cohort should be similar if they, e.g., refer to events before a fixed age.


## Comparison of NATSAL-1990 with NATSAL-2000

- Homosexual experience before 1990 [High sensitivity]:
- Men: $5.0 \%$ (1990) vs $8.5 \%$ (2000)
- Women: 3.5\% (1990) vs $6.7 \%$ (2000)
- Heterosexual intercourse before 16 years [Medium sensitivity]:
- Men: $24.7 \%$ (1990) vs $27.5 \%$ (2000)
- Women: 12.9\% (1990) vs $18.2 \%$ (2000)

From these type of questions. estimate odds ratios (ORs) for change in bias

- High sensitivity: Men $1.80(1.46,2.21) ;$ Women 1.99(1.62,2.46)
- Medium sensitivity: Men 1.11(1.01,1.21); Women 1.19(1.10,1.29)


## Comparison of NATSAL-1990 with NATSAL-2000

Homosexual partners, past 5 years

- Men: $1.5 \%$ (1990) vs $2.6 \%$ (2000) $\rightarrow$ OR: $1.75(1.29,2.36)$
-:
- Women: 0.8\% (1990) vs 2.6\% (2000) $\rightarrow$ OR: 3.43(2.42,4.87)
-:


## Comparison of NATSAL-1990 with NATSAL-2000

Change in bias results

- High sensitivity OR: Men 1.80; Women 1.99

Homosexual partners, past 5 years

- Men: $1.5 \%$ (1990) vs $2.6 \%$ (2000) $\rightarrow$ OR: $1.75(1.29,2.36)$
- Minimum established change: $1.29 / 1.80=0.72$
- Women: $0.8 \%$ (1990) vs $2.6 \%$ (2000) $\rightarrow$ OR: 3.43(2.42,4.87)
- Minimum established change: $2.42 / 1.99=1.22$


## Concluding Remarks

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- The collection and use of auxiliary information which is directly linked to missing or informatively collected data should be sought in such modelling efforts.
- The type of information and appropriate model is likely to be application specific.
- Caution is still strongly advised.

